

Tomographic Microwave Diversity Image Reconstruction Employing Unitary Compression

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Abstract—Data compression through a unitary transform is utilized in tomographic microwave diversity image reconstruction in order to reduce the dimensionality and to extract the features in the data space. The unitary compression is derived by minimizing the mean-square error (MSE) and the unitary transform is made of eigenvectors of the data's covariance, regarded to be a Karhunen–Loève transform. Tomographic microwave frequency-swept imaging was developed using a unique target-derived reference technique to access the three dimensional Fourier space of the scatterer and an image reconstruction algorithm based on the projection slice theorem. It is shown that: centimeter resolution of a complex object can be preserved even when half of the data set is compressed; and that the reconstructed image remains identifiable by a human observer even when 2/3 of the data set is compressed.

I. INTRODUCTION

TOMOGRAPHIC microwave diversity imaging was developed and has been in use at the Electro-Optics & Microwave-Optics Laboratory at the University of Pennsylvania for over a decade. It demonstrates the feasibility of broad-band projective imaging, with near-optical resolution, of complex shaped scattering object in the 2–26.5 GHz range [1]–[4]. The basic working principle of tomographic microwave diversity imaging relies on the Bojarski identity of inverse scattering theory [7], [8]. Based on this principle one can access a 2-D slice of the 3-D Fourier space of the scattering object by measuring the scattered field from an object that can be rotated in azimuth relative to a fixed broad-band transmitter/receiver (T/R) assembly and by digitally sweeping the transmitted signal over a bandwidth of several GHz for each rotation angle and coherently detecting the back-scattered wave field. The Fourier space data acquired in this fashion, is made of frequency response echos which are measured at different aspect views of the object which are equally spaced over a prescribed angular window. Each echo is quantized over the sweeping frequency spectrum. If these frequency response echos are considered as samples of a stochastic process, they are mutually correlated and have some redundancy.

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In the practical applications of tomographic microwave diversity imaging, there is an emerging problem of how to deal with the vast amount of 2-D data acquired from an imaging radar. For instance, target identification methods that form an associative memory based on neural networks [4] require that the data set be reduced. Similar problems appear in telecommunication and communication networks where the capacity of the transmission channel is limited and signal (data/image/voice) to be transmitted have to be compressed before the transmission in order to achieve fast and least distorted transmission. For all these data compression problems, the key point is how to find a signal representation identifying the redundancy of the signal to be compressed. Although no representation is best in all senses, the unitary transform has been widely used to serve as one of the signal representation in a lot of applications such as image processing [5], speech processing [6], feature selection in pattern recognition [17], and analysis and design of communication systems [11]. Unitary compression is a technique that applies the unitary transform to data compression, which allows the redundancy of the data to be identified and thus removed. The procedure of applying the unitary transform to the data compression is described as follows: the measured Fourier data is unitarily transformed to a domain in which coordinates are de-correlated, and redundant ones are identified and removed; the new Fourier space data is then obtained by the inverse unitary transform while maintaining a tolerable fixed distortion.

Based on the criterion of minimizing the mean-square error, a unitary transform, called Karhunen–Loève transform (KLT), can be given, which will make the vectors in the data space orthogonal with respect to each other. As pointed out by Watanabe [13], the Karhunen–Loève expansion enables the extraction from a given set of data those features which are the most significant, i.e., which bear the greatest amount of information. To do this one works with the covariance matrix averaged over the whole data set and diagonalizes it. The data features are then the eigenvectors and the dominance or importance of these features are the eigenvalues. While a subset of the data with respect to smaller eigenvalues of the data's covariance matrix is defined as redundant ones, the mean-square error, which is derived as a summation of eigenvalues of

axis of the rotation, i.e., the shape estimation of the object shown in Fig. 2(a). Resolution of the reconstructed image is estimated to be about 2 cm from observed details such as the engines and tail section in this image.

III. KARHUNEN-LOÉVE TRANSFORM

The unitary transforms, defined in a linear inner-product space in a complex field which is also called unitary space, are of complex forms of the orthogonal transforms. The Karhunen-Loève transform is one of them and will be outlined for review and to provide background for the notation used in this paper. It is assumed that a given X is a second-order discrete stochastic process and that its covariance is denoted by a 2-D matrix Σ_x defined by

$$\Sigma_x = E[(x - m_x)(x - m_x)^*] \quad (3)$$

where $E[\cdot]$ denotes the statistical expectation; m_x is the 1-D mean vector of X ; and $*$ is complex transpose. Since X is a complex process, Σ_x is Hermitian (i.e., complex conjugate with real diagonal elements). We are interested only in the truncated version of the Karhunen-Loève transform because it has been shown [12] to yield the best approximation of a stochastic process of all N -dimensional approximations.

Let $T^* = [\phi_1, \dots, \phi_N]$ be a complete, unitary basis such that X can be expanded in the sense of the mean-square convergence:

$$X \doteq \sum_{i=1}^N y^{(i)} \phi_i \quad (4)$$

where $y^{(i)}$ are coordinates vector, and the notation \doteq represents the mean-square equivalence, i.e.,

$$E \left[\left(X - \sum_{i=1}^N y^{(i)} \phi_i \right)^2 \right] = 0. \quad (5)$$

Owing to the unitary characteristics of the basis vector T , the coordinates $y^{(i)}$ can be obtained by

$$y^{(i)} = \langle X, \phi_i \rangle \quad (6)$$

where $\langle \cdot, \cdot \rangle$ denotes a complex inner-product for the discrete case. The expansion given by (4) is regarded as the Karhunen-Loève expansion when the basis vector T is chosen such that the coordinates are unitary in the probabilistic sense:

$$E[(y^{(i)} - m_y^{(i)})(y^{(j)} - m_y^{(j)})^*] = \lambda_i \delta_{ij} \quad (7)$$

where δ_{ij} is the Kronecker delta function; λ_i , an eigenvalue of X 's covariance, is given by

$$\Sigma_x \phi_i = \lambda_i \phi_i \quad (8)$$

where Σ_x is a Hermitian and non-negative definite matrix, and ϕ_i are eigenvectors of Σ_x . The Karhunen-Loève transform is therefore considered a bi-unitary transform because both the coordinates and the basis vectors are unitary. It is shown that the Karhunen-Loève transform is made of the eigenvectors of the data's covariance.

If we denote $Y^* = [y^{(1)}, \dots, y^{(N)}]$, then (4) can be expressed in matrix format:

$$X = T^* Y \quad (9)$$

Since $\{\phi_i, i = 1, \dots, N\}$ are unitary with respect to each other, i.e.,

$$\phi_i \phi_j = \begin{cases} 1 & i = j \\ 0 & i \neq j \end{cases}, \quad (10)$$

then

$$TT^* = I, \quad (11)$$

where I is the identity matrix, and (11) can be rewritten by

$$T^{-1} = T^*. \quad (12)$$

Hence, the inverse of a unitary matrix is the complex transpose of the matrix. (9) can, therefore, be rewritten by multiplying T using (11)

$$Y = TX. \quad (13)$$

The covariance of Y is then given by

$$\Sigma_y = T \Sigma_x T^{-1} = T \Sigma_x T^*. \quad (14)$$

From linear algebra, it is known that a Hermitian matrix can be diagonalized by a unitary matrix consisting of eigenvectors of the Hermitian matrix, with the eigenvalues of the Hermitian matrix as its diagonal elements [14]. Since the covariance matrix Σ_x is Hermitian and T is consisting of eigenvectors of the matrix Σ_x , the covariance matrix Σ_y given by (14) can be diagonalized with the eigenvalues of Σ_x as its diagonal elements. Thus, (14) can be rewritten by

$$\begin{aligned} \Sigma_y &= T \Sigma_x T^{-1} \\ &= T \Sigma_x T^* \\ &= \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_N \end{bmatrix}. \end{aligned} \quad (15)$$

This result implies that the coordinates matrix Y is unitary because its covariance matrix is diagonal. Since the unitary transform is an inner-product invariant, norm invariant transform, the variance is not changed after the Karhunen-Loève transform. The variance of X can be expressed by the diagonal terms of Y

$$\begin{aligned} \text{Var}(X) &= \text{Var}(Y) \\ &= \sum_{i=1}^N \lambda_i \end{aligned} \quad (16)$$

where λ_i are the eigenvalues of Σ_x .

IV. IMAGE RECONSTRUCTION EMPLOYING UNITARY COMPRESSION

A. Theoretical Foundation

The image reconstructed from a sufficient quantity of acquired Fourier space data is displayed in Fig. 2. However, the acquired Fourier space data is correlated and there exists redundancy in terms of correlation. The unitary compression is a technique in which the mean-square error due to eliminating the redundant data is minimized through a unitary transform. This unitary transform allows us to extract the most significant features of the acquired data and to compress the data sets so that their redundancy can be removed. The procedure of applying the unitary transform to the image reconstruction or image transmission is depicted in Fig. 3. The Fourier space data X is unitarily transformed to a domain in which coordinates $y^{(i)}$ are de-correlated, and redundant ones are identified and removed. The new Fourier space data \hat{X} is obtained by the inverse unitary transform T^{-1} while maintaining a tolerable fixed distortion between X and \hat{X} . Thus, \hat{X} is considered an effective representation containing the most significant information and is used to reconstruct images.

According to Fig. 3, for the acquired data vector X , we obtain

$$Y = TX; \quad (17)$$

where

$$\begin{aligned} X^* &= [x^{(1)}, \dots, x^{(N)}]; \\ Y^* &= [y^{(1)}, \dots, y^{(N)}]; \end{aligned}$$

and

$$T^* = [\psi_1, \dots, \psi_N].$$

From (17) and it follows that $TT^* = I$, and hence

$$\begin{aligned} X &= T^* Y \\ &= \sum_{i=1}^N y^{(i)} \psi_i. \end{aligned} \quad (18)$$

where X can be referred to as an expansion in terms of a complete unitary bases $\psi_i (i = 1, \dots, N)$ and $y^{(i)}$ are the coordinates. If we wish to retain a subset $\{y^{(1)}, \dots, y^{(M)}\}$ of $Y (M \leq N)$ and yet estimate X . This can be done by replacing remaining $N - M$ components of Y by pre-selected constants vector $b^{(i)} (i = M + 1, \dots, N)$ to obtain

$$\hat{X} = \sum_{i=1}^M y^{(i)} \psi_i + \sum_{i=M+1}^N b^{(i)} \psi_i. \quad (19)$$

where \hat{X} denotes the estimate of X . The error ΔX introduced by neglecting the $N - M$ terms can be represented

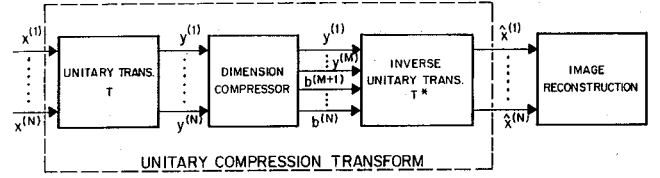


Fig. 3. Block diagram of the image reconstruction procedure with unitary compression transform.

as

$$\begin{aligned} \Delta X &= X - \hat{X} \\ &= X - \sum_{i=1}^M y^{(i)} \psi_i - \sum_{i=M+1}^N b^{(i)} \psi_i \\ &= \sum_{i=1}^N y^{(i)} \psi_i - \sum_{i=1}^M y^{(i)} \psi_i - \sum_{i=M+1}^N b^{(i)} \psi_i \\ &= \sum_{i=M+1}^N (y^{(i)} - b^{(i)}) \psi_i. \end{aligned} \quad (20)$$

Thus the mean-square error ϵ is given by

$$\epsilon = E[\|\Delta X\|^2] = E[\Delta X \cdot \Delta X^*] \quad (21)$$

where $\|\cdot\|$ is a norm notation. Substitution of (20) in (21) leads to

$$\begin{aligned} \epsilon &= E \left[\sum_{i=M+1}^N \sum_{j=M+1}^N (y^{(i)} - b^{(i)}) (y^{(j)} - b^{(j)})^* \psi_i \psi_j^* \right] \\ &= \sum_{i=M+1}^N E[(y^{(i)} - b^{(i)})^2]; \end{aligned} \quad (22)$$

where the unitary property of ψ_i is used to simplify the result in (22). In order to minimize the ϵ by choosing both $b^{(i)}$ and ψ_i , we take the derivatives of ϵ with respect to $b^{(i)}$ and ψ_i and set them to zero, that is

$$\frac{\partial \epsilon}{\partial b^{(i)}} = -2\{E[y^{(i)}] - b^{(i)}\} = 0 \quad (23)$$

which yields

$$\begin{aligned} b^{(i)} &= E[y^{(i)}] \\ &= m_y^{(i)} \\ &= \psi_i^* m_x^{(i)} \end{aligned} \quad (24)$$

where $m_x^{(i)}$ and $m_y^{(i)}$ represents the means of X and Y , respectively. Hence, the mean-square error ϵ can be rewritten after substituting (24) into (22):

$$\begin{aligned} \epsilon &= \sum_{i=M+1}^N \psi_i E[(x - m_x)(x - m_x)^*] \psi_i^* \\ &= \sum_{i=M+1}^N \psi_i^* \Sigma_x \psi_i; \end{aligned} \quad (25)$$

where

$$\Sigma_x = E[(x - m_x)(x - m_x)^*]$$

is the covariance matrix of X .

To obtain the optimal $\psi_i (i = 1, \dots, N)$ we must not only minimize ϵ with respect to ψ_i , but also satisfy the unitary constraint:

$$\psi_i \psi_j = \delta_{ij};$$

or

$$TT^* = I.$$

Thus we use the method of Lagrange multipliers and minimize

$$\begin{aligned} \tilde{\epsilon} &= \epsilon - \sum_{i=M+1}^N \theta_i [\psi_i^* \psi_i - 1] \\ &= \sum_{i=M+1}^N \psi_i^* \Sigma_x \psi_i - \theta_i [\psi_i^* \psi_i - 1]; \end{aligned} \quad (26)$$

with respect to ψ_i , where θ_i denotes the Lagrange multiplier. By taking the derivatives with respect to ψ_i , we have

$$\frac{\partial \epsilon}{\partial \psi_i} = \nabla_{\psi_i} [\psi_i^* \Sigma_x \psi_i] - 2\theta_i \nabla_{\psi_i} [\psi_i^* \psi_i] \quad (27)$$

where ∇ denotes a vector derivative operator. The following can be given from the differentiability of the matrix:

$$\nabla_{\psi_i} [\psi_i^* \Sigma_x \psi_i] = 2\Sigma_x \psi_i; \quad (28)$$

and

$$\nabla_{\psi_i} [\psi_i^* \psi_i] = 2\psi_i. \quad (29)$$

Thus, from (27) we have

$$\frac{\partial \epsilon}{\partial \psi_i} = 2\Sigma_x \psi_i - 2\theta_i \psi_i = 0; \quad (30)$$

which yields

$$\Sigma_x \psi_i = \theta_i \psi_i, \quad (31)$$

in which ψ_i are the eigenvectors of the covariance matrix Σ_x , i.e., $\psi_i = \phi_i (i = 1, \dots, N)$, and θ_i are the eigenvalues, i.e., $\theta_i = \lambda_i (i = 1, \dots, N)$. As previously discussed, the unitary transform T , consisting of eigenvectors of the covariance matrix, is called Karhunen-Lo  ve transform. It has been shown that Y is unitary (i.e., uncorrelated), which makes it useful in identification where the unique features of the data can be extracted from the unitary data sets based on *Principal Component Analysis* [17] in which the redundancy of the data can be discerned [18]. The mean-square error can be given by substituting (31) in (25):

$$\begin{aligned} \epsilon &= \sum_{i=M+1}^N \phi_i^* (\lambda_i \phi_i) \\ &= \sum_{i=M+1}^N \lambda_i. \end{aligned} \quad (32)$$

The mean-square error caused by discarding the $N - M$ redundant vectors is the summation of the $N - M$ eigenvalues of the covariance matrix Σ_x . This result indicates that the effectiveness of a transform component $y^{(i)}$ for representing the data vector X , is determined by its corresponding eigenvalue. If a component $y^{(i)}$ is deleted, then the mean-square error increases by λ_i , the corresponding eigenvalue. Thus the set of $y^{(i)}$ with the M largest eigenvalues should be selected, and the remaining $y^{(i)}$ discarded since they can be replaced by the constants $b^{(i)}$, $i = M + 1, \dots, N$. It is, therefore, obvious that in order to minimize the mean-square error while maximizing the possibility of compressing the data (i.e., maximizing M), the vectors chosen to be discarded should be these with smaller eigenvalues. Based on the above analysis, two important conclusions are given as follows: 1) The KLT is the optimal transform for unitary compression with respect to the mean-square error criterion. Since the most significant features about the data are dominated by only some transform components, removing the redundancy of the data will be possible; 2) Since $Y = TX$, the transform domain covariance Σ_y is expressed by

$$\Sigma_y = T\Sigma_x T^{-1} = T\Sigma_x T^* \quad (33)$$

where T consists of the eigenvectors of Σ_x , which yields

$$\Sigma_y = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_N \end{bmatrix}. \quad (34)$$

This implies that the vector Y is unitary and uncorrelated.

Since the eigenvalues are the main diagonal terms of Σ_y , they correspond to the variances of the transform components $y^{(i)}$, $i = 1, \dots, N$. Thus a logical criterion for selecting transform components is to retain the set of components with the M largest variance, while the remaining $(N - M)$ components are discarded. This process of components selection is referred to as the *variance criterion* method. Since the total variance of X is the sum of all eigenvalues of Σ_x , the variance-loss-ratio is defined by

$$\text{VLR} = \frac{\text{mean-square error}}{\text{total variance}} \times 100\%; \quad (35)$$

where the mean-square error is caused by eliminating some vectors. The compression-ratio is defined as

$$\text{CR} = \frac{\text{no. of vectors eliminated}}{\text{total no. of vectors}} \times 100\%. \quad (36)$$

These two ratios are used as the important performance indexes to test the effectiveness of the unitary compression.

B. Experimental Results

The experimental setup was as described in Section II, and the experimental conditions and parameters are same as these given previously. The test object, a metallized

100:1 scale model of B-52 aircraft, is mounted on a computer-controlled position situated in the anechoic chamber. The received Fourier space slice data is composed of 128 views spanning 90 degrees rotation of the test object (i.e., 128 equally spaced views spanning a 90 degrees angular window), and 201 sampling points for each view covering a 2–17 GHz frequency bandwidth. Therefore, the measured scattering data set for the target is a 128×201 matrix representing 128 aspect views, each of which has 201 sampling points. The covariance representing the correlations between aspect views, can be formed for the 128 aspect views. Since the scattering data are of complex numbers, the covariance can be given by

$$\Sigma_x^{ij} = \frac{1}{N_{\text{samp}} - 1} \sum_{k=1}^{N_{\text{samp}}} (\mathbf{x}_k^{(i)} - \mathbf{m}_x^{(i)}) (\mathbf{x}_k^{(j)} - \mathbf{m}_x^{(j)})^*, \quad (37)$$

where * denotes the complex transpose, N_{samp} denotes the number of the sampling points, and $\mathbf{m}_x^{(i)}$ is the mean of $\mathbf{x}^{(i)}$, which is computed as

$$\mathbf{m}_x^{(i)} = \frac{1}{N_{\text{samp}}} \sum_{k=1}^{N_{\text{samp}}} \mathbf{x}_k^{(i)} \quad (i = 1, \dots, N_{\text{view}}) \quad (38)$$

where N_{view} denotes the number of the aspect views. The covariance matrix Σ_x^{ij} is Hermitian and takes a maximum value when $i = j$. Both eigenvalues and eigenvectors of the covariance matrix are computed by using IMSL, a popular Fortran library for mathematical applications [19]. The Fourier space slice data received is transformed by the KLT which consists of the eigenvectors of the data's covariance and the data in transform domain are unitary, or uncorrelated with respect to each other. The eigenvalues solved by

$$\Sigma_x \phi_i = \lambda_i \phi_i \quad (39)$$

are illustrated in Fig. 4, where the eigenvalues, because of Hermitian nature of the covariance matrix, are real valued and can be ranked by magnitude. Since the eigenvalues are the main diagonal terms of Σ_y , they correspond to the variance of the transform component $\mathbf{y}^{(i)}$. Thus, Fig. 4 will be referred to as the variance distribution. The area under the variance distribution curve for a given number of transform components is an indication of the amount of energy contained in those components. It is apparent that almost all of the signal energy (i.e., area under the curve) is packed into first half or one-third KLT components. Therefore, the most important information needed for image reconstruction is concentrated in only a few dominating aspect views and there is a great deal of redundancy in the raw data. According to *variance criterion*, these transform components corresponding to the larger eigenvalues contain the most significant features and should be kept, while the remaining with respect to smaller eigenvalues are considered as the redundant ones and discarded. The mean-square error caused by eliminating the redundant data can be minimized by discarding these aspects views with smaller associated eigenvalues.

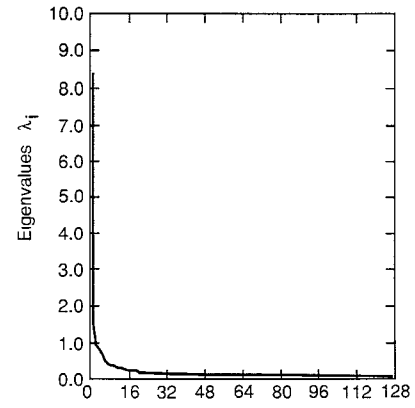


Fig. 4. Ordered eigenvalues of data's covariance.

TABLE I
VARIANCE-LOSS-RATIO (VLR)
VERSUS COMPRESSION RATIO (CR)

CR (%)	VLR (%)
$\frac{10}{128} \approx 7.8\%$.01%
$\frac{50}{128} \approx 39\%$	1%
$\frac{64}{128} \approx 50\%$.15%
$\frac{100}{128} \approx 78\%$	4%
$\frac{120}{128} \approx 94\%$	18.6%

The relation between VLR and CR (both defined above) is demonstrated in Table I, where it is seen that, even when 64 vectors (half of the total number) are eliminated, the loss (i.e., mean-square error) is small, $\text{VLR} = \frac{1}{1000}$. If the signal energy is essentially contained in 43 of the 128 KLT components, a 3:1 data compression is considered. For a 3:1 compression, 43 KLT components corresponding to the largest diagonal terms of Σ_x are selected. Such a compressed unitary data set is denoted by $\hat{\mathbf{Y}}$. An inverse KLT which, because of its unitary, is a complex transpose of T , is employed to reconstruct the image. Thus, the data to be used for image reconstruction are $\hat{\mathbf{X}} = T^* \hat{\mathbf{Y}}$. The images reconstructed by employing unitary compression are presented in Figs. 5 and 6. As illustrated in Fig. 5 where half of the coordinates $\mathbf{y}^{(i)}$ (64) are discarded, the reconstructed image is still legible and has same centimeter resolution as the original one shown in Fig. 1(a). Fig. 6 shows the reconstructed image when 2/3 of the received data have been discarded, i.e., a 3:1 data compression is achieved, in order to demonstrate that the centimeter resolution can be preserved. It is apparently shown in Fig. 6 that two engines on the wing can still be recognized, although there is some blurring. In conclusion, the unitary compression is an effective technique for dimension reduction in image reconstruction. It presents the statistical optimal solution for the data compression and feature extraction in the sense of minimizing the mean-square error. The *variance criterion* has

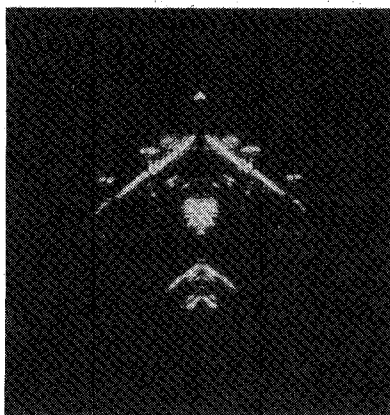


Fig. 5. Result of image with UCT when $CR = 1/2$ (half of the data are compressed).

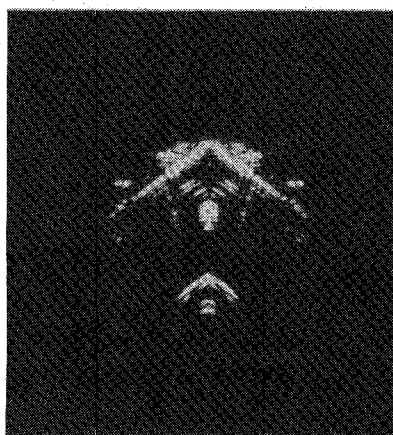


Fig. 6. Result of image with UCT when $CR = 2/3$ (two thirds' of the data are compressed).

demonstrated the capability of identifying the redundancy of the correlated data sets.

V. CONCLUSION

This work was initiated to study the feasibility of applying unitary compression to tomographic microwave diversity image reconstruction. It was found that a unitary transform, the Karhunen-Loève transform, can provide a good representation of the data so that the redundancy of the data can be identified and thus allows for removal. This representation is based on minimizing the mean-square error, which leads to a statistical optimal solution. The mean-square error caused by discarding the redundant data is derived as the sum of those eigenvalues corresponding to the discarded data. Therefore, the eigenvalues of the data's covariance matrix are used as features to identify the redundancy of the data. According to the *variance criterion*, the discarded data, which is considered redundant, should be these with the smaller eigenvalues. The reconstructed images employing the unitary compression show that image quality is basically preserved after half of the total data has been selectively eliminated; and one can still identify most of details of the test model even when two-third's data are discarded.

Although KLT is the optimal solution for unitary data compression, there are no fast computational algorithms since it involves solving the eigenvalues and eigenvectors of the data's covariance matrix. One is tempted, therefore, to employ suboptimal unitary transforms that possess fast computational algorithms. Discrete cosine transform (DCT) will be considered as an alternative unitary transformation for unitary compression. The performance evaluation of the effectiveness of unitary transforms will be studied in terms of rate-distortion function as a performance measure.

Targets can be identified either by reconstructing images with sufficient resolution to be recognized by a human observer, or by extracting features of the target for automated machine recognition. In future work, we will explore the use of unitary compression transform in neuromorphic target recognition. The unitary compression transform is used then as a feature detector in which the principal component can be extracted from the eigenvalues of the data's covariance matrix. The feature extraction based on the principal component method is by self-organization as shown in Oja's work [19]. Moreover, the unitary compression should be more effective when it is applied to target recognition, since it is expected that there is more redundancy in the data for this task than when the data are used for image reconstruction.

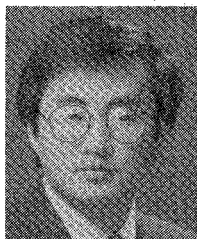
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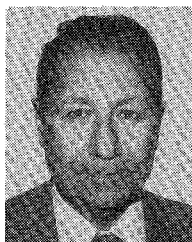
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